Random graphs with hidden color

Bo Söderberg*

Complex Systems Division, Department of Theoretical Physics, Lund University, Sölvegatan 14A, S-22362 Lund, Sweden
(Received 23 April 2003; published 15 July 2003)

We propose and investigate a unifying class of sparse random graph models, based on a *hidden coloring* of edge-vertex incidences, extending an existing approach, random graphs with a given degree distribution, in a way that admits a nontrivial correlation structure in the resulting graphs. The approach unifies a number of existing random graph ensembles within a common general formalism, and allows for the analytic calculation of observable graph characteristics. In particular, generating function techniques are used to derive the size distribution of connected components (clusters) as well as the location of the percolation threshold where a giant component appears.

INTRODUCTION

There is a growing interest in complex networks, in the physics community as well as in other sciences, partly due to an increased availability of data on real-world networks. This is reflected in a rapidly increasing number of models of random graphs [1–5] and dynamical random graphs [6–10], with varying degrees of generality.

This multitude of models calls for a unifying formalism, including more specific models as special cases, while allowing for the calculation of observable characteristics that can be compared to those of real networks. Dynamical models are interesting in their own rights, but the dynamics is seldom directly observable in real-world networks, and we will focus on static ensembles of random graphs, irrespective of whether they result from a dynamical process or not.

Specifically, we will consider models of simple, undirected graphs that are *sparse* (the edge count grows linearly with the node count N) and *truly random* (having no underlying regular structure). The *classic random graph* in its sparse version is of this type [1,11,12], where each of the $N(N-1)/2$ possible edges is independently and randomly realized with a fixed probability $p=c/N$. It has a Poissonian asymptotic degree (connectivity) distribution with average c and a percolation threshold at $c=1$. It fails, however, to describe most real-world networks.

Instead, we turn to two of the more general approaches, based on slightly different philosophies. One, to be referred to as DRG (degree-driven random graph), amounts to choosing a random member from the set of simple labeled graphs with a given arbitrary degree distribution [2,13,14]. The other is inhomogeneous random graph [4], IRG, where the classic model is generalized by randomly coloring vertices according to a color distribution $\{r_i\}$, and realizing edges independently with color-dependent probabilities c_{ij}/N. Both yield analytically tractable models displaying well-defined percolation thresholds and degree distributions, both include a number of more specific models, and both have limitations: DRG fails to produce nontrivial edge correlations, as seen in the factorization of the combined degree

distribution of connected vertex pairs [7]; in IRG, the resulting degree distribution is limited to a Poissonian mix [4].

These approaches are not unrelated: The restriction of DRG to degree distributions in the form of a Poissonian mix is in fact asymptotically equivalent to the restriction of IRG to a rank-one c matrix, $c_{ab}=C_a C_b$ (exhibiting DRG's lack of correlations); this common subset contains the classic model [4].

BASIC IDEA

By combining the philosophies of DRG and IRG, a more general class of analytically tractable sparse random graph models can be constructed. This unifying approach, to be referred to as CDRG (for colored DRG), contains IRG and DRG as particular subsets, and is defined as a direct extension of DRG by assigning a *hidden color* to each vertex connection (a half edge or *stub*). As a result, each edge will be associated with a pair of colors, one for each end point. We then consider a given distribution $\{p_{\mathbf{m}}\}$ of *colored degrees* $\mathbf{m}=(m_1,\ldots,m_K)$, where for each vertex its number m_k of stubs of each color k is accounted for, and allow the edge distribution to be color-sensitive by specifying also the distribution of edge color pairs. The resulting ensemble of stub-colored graphs yields, if the coloring is considered *unobservable*, a well-defined graph ensemble. The coloring thus can be thought of as a set of *hidden variables*, the purpose of which is to induce correlations in the resulting graphs.

Below, we will discuss the definition and implementation of CDRG models, derive the asymptotic cluster size distribution yielding equations for the percolation threshold, and identify the subsets corresponding to DRG (trivial) and IRG (less trivial).

ASYMPTOTIC MODEL SPECIFICATION

A particular asymptotic CDRG model is defined by specifying the following features.

(1) A definite color space, say $\{1,2,\ldots,K\}$.

(2) An asymptotic colored degree distribution (CDD), $p_{\mathbf{m}}$, defining the relative frequencies of vertices with different colored degrees $\mathbf{m}=(m_1,\ldots,m_K)$, where m_a is the

number of a-colored stubs of the vertex. We will assume here that all its moments, $\langle m_a \rangle \equiv \Sigma_{\mathbf{m}} p_{\mathbf{m}} m_a$, $\langle m_a m_b \rangle$, etc., are defined.

(3) A symmetric, non-negative $K \times K$ *color preference matrix* **T**, controlling the relative abundance $\sim \langle m_a \rangle T_{ab} \langle m_b \rangle$ of edges between different color pairs a, b. It must satisfy

$$\sum_{b=1}^{K} T_{ab} \langle m_b \rangle = 1. \tag{1}$$

Note that the total degree of a vertex is simply the sum of its colored degree components; the usual degree distribution is thus also fixed, and amounts to $p_m = \Sigma_{\mathbf{m}} \delta(m, \Sigma_a m_a) p_{\mathbf{m}}$.

TRUNCATION TO FINITE N

We want to implement such an asymptotic model with a specific N. This can be done, e.g., by transforming the CDD into a definite colored degree sequence, as described by the number of vertices $N_{\mathbf{m}} \approx N p_{\mathbf{m}}$ with colored degree $\mathbf{m}$, subject to obvious constraints such as $m < N$, $\Sigma_{\mathbf{m}} N_{\mathbf{m}} = N$, and $\Sigma m N_{\mathbf{m}}$ is even. Similarly, the matrix **T** is used to determine the number of edges with color-pair ab as $n_{ab} \approx N \langle m_a \rangle T_{ab} \langle m_b \rangle$. Note that each ab edge is counted twice, as ab and as ba, so the diagonal elements n_{aa} must be even. The number of edge end points (*butts*) with color a becomes $n_a = \Sigma_b n_{ab} \approx N \langle m_a \rangle \Sigma_b T_{ab} \langle m_b \rangle$, and care must be taken so that this matches the corresponding number of stubs, $\Sigma_{\mathbf{m}} m_a N_{\mathbf{m}} \approx N \langle m_a \rangle$—thus constraint (1) on **T**.

This yields a pool of vertices with definite colored degrees and a pool of edges with definite color pairs, all to be considered distinguishable. The set of distinct ways to combine these into a simple graph with color-matching between butts and stubs defines a set of colored graphs. By drawing a random member from this set and neglecting the coloring, the desired truncated CDRG ensemble results.

IMPLEMENTATION IN PRACTICE

When it comes to the practical task of generating random graphs from this ensemble, the difficult step is that of picking a random member from the set of colored graphs, consistent with definite $N_{\mathbf{m}}$ and n_{ab}. A random stub-pairing method for DRG [2] can be extended to the case of colored stubs as follows.

(1) For each color a, make a complete random assignment between the n_a butts of color a and the n_a matching stubs, to determine which butt should attach to which stub.

(2) While the resulting graph is not simple, repeat step (1).

Alternatively, the implementation could be done in a fully stochastic manner, where an extra initial step is to draw N colored degrees independently from $p_{\mathbf{m}}$, and a pool of edges from $q_{ab} = \langle m_a \rangle T_{ab} \langle m_b \rangle / \Sigma_c \langle m_c \rangle$, subject to matching counts of stubs and butts of each color. In the thermodynamic limit, the result would be equivalent. Such a method would be more in line with the identification of CDRG with the Feynman graphs of zero-dimensional multicomponent field theories, in analogy with the relation between DRG

models and zero-dimensional scalar field theories [15]

Of course, either generation method is feasible only if the probability of obtaining a simple graph in each pairing attempt is not too small. This probability is asymptotically calculable.

PAIRING EFFICIENCY

A completely random pairing without the restriction that the resulting graph be simple yields an ensemble of *multigraphs*, i.e., possibly nonsimple graphs where *loops* (cycles of length 1) and/or *multiple edges* are allowed. The efficiency of the above method depends on the probability to obtain a simple graph, which in turn depends on the abundance of loops and multiple edges. In a sparse graph, the probability for an edge between a given pair of nodes scales as $1/N$, so we expect a finite number both of double edges (a factor of N^2 for the choice of a node pair and $1/N^2$ for two edges) and of loops (N for the choice of node and $1/N$ for the edge making a loop).

In fact, we can compute the asymptotically expected number of loops and double edges in a random pairing to leading order:

Loops. For a single vertex with colored degree **m**, the probability that two of its stubs will be connected is given by $\Sigma_{ab}(m_a m_b - m_a \delta_{ab}) T_{ab}/2N$. Averaging over **m** and summing over the node choice yields the expected number of loops as $\alpha = \Sigma_{ab} M_{ab} T_{ab}/2$, i.e.,

$$\alpha = \tfrac{1}{2} \mathrm{Tr}(\mathbf{TM}), \tag{2}$$

where $\mathbf{M} = \{M_{ab}\}$ stands for the matrix of moments $\langle m_a m_b - m_a \delta_{ab} \rangle$.

Double edges. Similarly, for an arbitrary pair of nodes with colored degrees $\mathbf{m}, \mathbf{m}'$, the probability of a double edge asymptotically amounts to $\Sigma_{abcd}(m_a m_b - m_a \delta_{ab})(m'_c m'_d - m'_c \delta_{cd}) T_{ac} T_{bd}/(2N^2)$. Averaging over $\mathbf{m}, \mathbf{m}'$ and summing over the choice of node pair yields the expected number of double edges as $\beta = \Sigma_{abcd} M_{ab} M_{cd} T_{ac} T_{bd}/4$, i.e.,

$$\beta = \tfrac{1}{4} \mathrm{Tr}(\mathbf{TM})^2, \tag{3}$$

while triple edges, etc., can be neglected altogether.

In a similar way, the asymptotically expected number of more general small subgraphs can be computed, which in particular enables the computation of the expectation of higher powers of the loop and double edge counts, resulting in the two counts asymptotically behaving as independent Poissonian random variables. Hence, the probability of obtaining a simple graph in the random pairing can be estimated as

$$\mathrm{Prob(simple)} \approx \exp(-\alpha - \beta). \tag{4}$$

As a result, an average of $\sim \exp(\alpha + \beta)$ pairing attempts is needed, rendering the method feasible for reasonably small $\alpha + \beta$; in other cases an alternative generation method will have to be employed, such as starting from an arbitrary col-

ored graph consistent with $N_{\mathbf{m}}, n_{ab}$, and applying a colored extension of a degree-preserving random rewiring algorithm suggested for DRG [16].

CONNECTED COMPONENT STATISTICS

The size distribution of the connected components (clusters) of a random graph can be probed by choosing an initial vertex at random and recursively following edges to new neighbors [14]. The sparsity of edges forces a finite set of revealed vertices to form a *tree* in the thermodynamic limit, since crosslinking is suppressed by factors of $1/N$. Hence, loops and double edges can be neglected to leading order, and the random color-matched pairing between stubs and butts reduces to a *random branching process* (branched polymer) based on the rules: (i) an edge emanating from a stub of color a ends in a stub of color b with probability $T_{ab}\langle m_b \rangle$; (ii) given the color b of a stub, it belongs to a vertex with colored degree $\mathbf{m}$ with probability $m_b p_{\mathbf{m}}/\langle m_b \rangle$.

The asymptotic random branching process is conveniently described in terms of a generating function $g(z) = \Sigma_n P_n z^n$ for the probability P_n that the connected component being revealed consists of n vertices. $g(z)$ can be expressed in terms of the corresponding generating functions $\mathbf{h}(z) = \{h_a(z)\}$ for the number of nodes in a branch starting from a stub of color a. $g(z)$ and $\mathbf{h}(z)$ satisfy the recursive relations

$$g(z) = z \sum_{\mathbf{m}} p_{\mathbf{m}} \prod_a h_a(z)^{m_a} \equiv z H(\mathbf{h}(z)), \quad (5a)$$

$$h_a(z) = z \sum_b T_{ab} \sum_{\mathbf{m}} p_{\mathbf{m}} m_b \prod_c h_c(z)^{m_c - \delta_{cb}}$$

$$\equiv z \sum_b T_{ab} \partial_b H(\mathbf{h}(z)), \quad (5b)$$

where $H(\mathbf{x}) = \Sigma_{\mathbf{m}} p_{\mathbf{m}} \mathbf{x}^{\mathbf{m}} \equiv \Sigma_{\mathbf{m}} p_{\mathbf{m}} \Pi_a x_a^{m_a}$ is the multivariate generating function for the CDD, while ∂_b stands for the derivative with respect to the bth argument of H. Equations (5) can be derived as follows. Equation (5a): The explicit factor of z accounts for the initial vertex, while the remainder consists of an average over the colored degree $\mathbf{m}$ of the initial vertex, of a factor $h_a(z)$ for each stub of color a, accounting for the contribution of the branch starting in that stub. Equation (5b): Starting from a stub of color a, the asymptotic probability that the other end of the attached edge has color b and is connected to a vertex having colored degree $\mathbf{m}$ is given by $T_{ab} p_{\mathbf{m}} m_b$; include a factor z for that vertex, and a factor $h_c(z)$ for each branch reached via one of its remaining $(m_c - \delta_{cb})$ stubs of color c.

PERCOLATION THRESHOLD

Of particular interest is the value of g for $z = 1$: naively, we expect $g(1) = h_a(1) = 1$, expressing the normalization of probability. Indeed, this defines a fixed point of recurrences (5), which however may be unstable. The stability can be analyzed by linearization of Eq. (5b) around $\mathbf{h}(1) = \mathbf{1}$, yield-

ing the Jacobian matrix $\mathbf{J}$ defined by

$$J_{ab} = \sum_c T_{ac} \partial_c \partial_b H(\mathbf{h})|_{\mathbf{h}=\mathbf{1}}, \quad (6)$$

which can be written as $\mathbf{J} = \mathbf{TM}$ [c.f. Eqs. (2) and (3)].

The point is that if an eigenvalue of $\mathbf{J}$ exceeds 1, the naive fixed point $\mathbf{h}(1) = \mathbf{1}$ turns unstable, signaling *supercriticality* of the branching process. In such a case, another fixed point will appear and take over as a stable solution with $h_a(1) < 1$, yielding $g(1) < 1$. Analogous phenomenona occur in the classic model as well as in IRG and DRG; the associated probability deficit $1 - g(1)$ is interpreted as the probability of hitting a *giant component* asymptotically containing a finite fraction $1 - g(1)$ of the vertices. This corresponds to a *percolating phase*; the percolation threshold is defined by the largest eigenvalue of $\mathbf{TM}$ being precisely 1.

INCLUSION OF OTHER MODELS

With a single color $K = 1$, CDRG trivially reduces to DRG, where a model is based on a given degree distribution $\{p_m\}$, while the preference matrix $\mathbf{T}$ reduces to a number, which by virtue of constraint (1) must equal $\langle m \rangle^{-1}$. Equations (5) reduce to the corresponding DRG equations

$$g(z) = z H(h(z)), \quad (7a)$$

$$h(z) = z \frac{H'(h(z))}{H'(1)}, \quad (7b)$$

with $H(x) \equiv \Sigma_m p_m x^m$ generating p_m. The percolating phase is defined by $J \equiv \langle m(m-1) \rangle / \langle m \rangle > 1$, yielding the well-known $\langle m(m-2) \rangle > 0$ [14].

The relation to IRG is less trivial: Assume the CDD to be in the form of a multi-Poissonian mix, i.e., $H(\mathbf{x}) = \Sigma_i r_i \exp[\Sigma_a C_{ia}(x_a - 1)]$. Define

$$g_i(z) \equiv z \exp\left(\sum_a C_{ia}(h_a(z) - 1) \right), \quad (8a)$$

$$c_{ij} \equiv \sum_{ab} C_{ia} T_{ab} C_{jb}, \quad (8b)$$

in terms of which Eqs. (5) reduce to

$$g(z) = \sum_i r_i g_i(z), \quad (9a)$$

$$g_i(z) = z \exp\left(\sum_j r_j c_{ij}(g_j(z) - 1) \right). \quad (9b)$$

Equations (9) exactly reproduce the result for $g(z)$ in an IRG model, with r_i taken as the probability of vertex color i and c_{ij}/N the probability of an edge between a pair of vertices with colors i, j [4].

Conversely, given an IRG model in terms of $\{r_i, c_{ij}\}$, one can always find $\{C_{ia}, T_{ab}\}$ satisfying Eq. (8b) such that $\Sigma_a C_{ia} = \Sigma_j c_{ij} r_j$.

It follows that CDRG contains also ensembles resulting from dynamical models such as randomly grown graphs [7] and dynamical random graphs with memory [8], which can be described in IRG [4], albeit at the cost of infinitely many colors.

CONCLUDING REMARKS

The above analysis shows that DRG and IRG can be unified into a more general class of random graph models, defined in terms of a hidden coloring of stubs and butts, with specified distributions of color-extended vertex degrees as well of edge color pairs. The purpose of the hidden coloring is to enable a nontrivial correlation structure in the resulting graphs.

This approach yields a general formalism for a large class of analytically tractable models on a given degree distribution, where local and global properties of the resulting graphs are calculable in the thermodynamic limit. Such a formalism also defines a suitable target for statistical model inference based on observed structural properties.

We have here assumed all moments of the degree distribution to exist, excluding, e.g., power behavior. The approach will be extended also to models with "fat tails." These are sensitive to the precise truncation method and will be treated elsewhere.

A more detailed investigation, addressing aspects and properties of CDRG models not treated in this Rapid Communication, is in progress and will be the subject of a forthcoming paper, as will the extension to directed graphs and to degree distributions with power tails.

Note added. Recently, the author was made aware of an article describing a related approach that can be seen as the restriction of CDRG to homogeneously colored vertices [17].

ACKNOWLEDGMENTS

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